

A probabilist's roadmap to the polygenic limit

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In front of a jury composed of:

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Charline SMADI	DR INRAE, Université Grenoble Alpes	Rapporteuse
Luis-Miguel CHEVIN	DR CNRS, CEFE	Examinateur
Joachim HERMISSON	PU Universität Wien	Examinateur
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Amandine VÉBER	DR CNRS, Université Paris Cité	Examinatrice
Amaury LAMBERT	PU École Normale Supérieure (ENS)	Directeur
Emmanuel SCHERTZER	PU Universität Wien	Directeur

The polygenic limit

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A **polygenic trait** is a measurable characteristic of an organism which is influenced by many genes.

Importance of the polygenic limit

- Interpreting the result of genomic data (Sella Barton 2019...)

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- Understanding the genetic architecture of diseases (Koch et al 2025+)

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- Understanding the genetic architecture of diseases (Koch et al 2025+)
- Survival of a population (Charlesworth 2013)
- Predicting long-term response of a population to a new environment (Hayward et al 2022...).

A tale of four queens

Selection (Darwin/Wallace 1859)

- There is a well-defined measurable quantity called the **fitness** of an organism which measures how well adapted an organism is to its current environment.



Image from Wikipedia

A tale of four queens

Selection (Darwin/Wallace 1859)

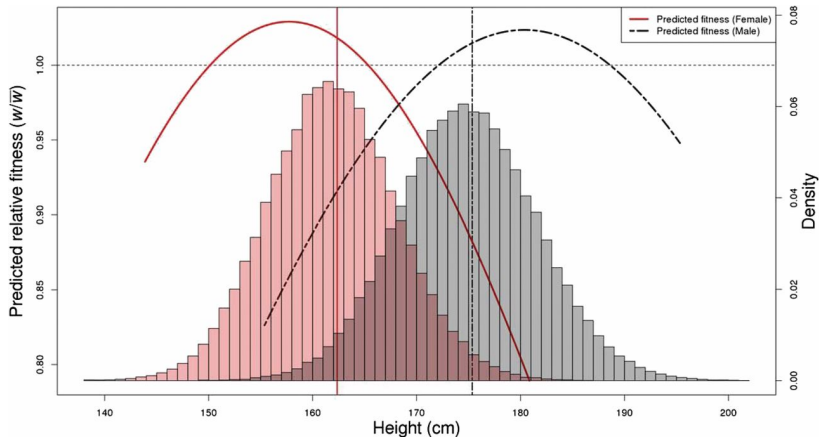
- There is a well-defined measurable quantity called the **fitness** of an organism which measures how well adapted an organism is to its current environment.
- The lower the fitness, the lower the chances to have healthy offspring.



Image from Wikipedia

A tale of four queens

Selection



From Sanjak et al, 2017

A tale of four queens

Mutation (Hugo de Vries 1901)

Heritable variability spontaneously appears within populations.



Image from Wikipedia

A tale of four queens

Recombination (Mendel 1865, Bateson, Saunders, Punnett 1905)

Offspring inherit some characteristics from each of their parents.



Image from Wikipedia

A tale of four queens

Recombination



From bartongroup.pages.ist.ac.at

A tale of four queens

Genetic drift (Hagedoorn, Hagedoorn-Vorstheuveel La Brand 1921, Wright 1929)

The evolution of a population is inherently random.



Image from Wikipedia

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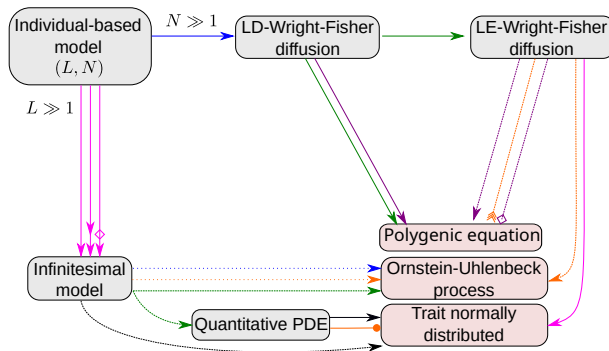
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- Migration, population structure, demography (Pólechova Barton 2015, Szép et al 2021 ...).

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- Sexual selection, homogamy (Surendranadh and Sachdeva, 2025)

Conclusion for mathematicians



Degree of proof

- Proven
- Discussed
-→ Not discussed

Type of proof

- Diffusion approximation
- Mixing
- Mean-field approximation
- Separation of timescales
- Central limit theorem

Complications

- Diploidy/Dominance
- Strong selection
- Epistasis
- Population structure/Demography

The individual-based model

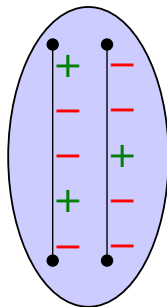
The individual-based model

The genome of an organism

The genome is structured as a pair of chromosomes with L loci.

Each locus has two alleles labelled $+$ and $-$.

A diploid genome can therefore be represented as an element of $(\{0, 1\}^L)^2$.



The individual-based model

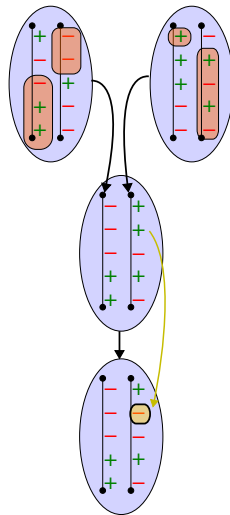
Inheritance

Recombination ♥:

Single uniform crossover

Mutation ♦:

locus ℓ has mutation probabilities $(\mu_{\ell}^{+}, \mu_{\ell}^{-})$



The individual-based model

Trait

Additive model: The **trait** of an organism with genome $G = (G_{\ell,i})_{\ell \in [L], i \in [2]} \in (\{0, 1\}^L)^2$ is

$$Z(G) := \sum_{\ell \in [L]} \alpha_{\ell} (G_{\ell,1} + G_{\ell,2})$$

with $\alpha_{\ell} \in \mathbb{R}_+$ the **additive effect on the trait at locus ℓ** .

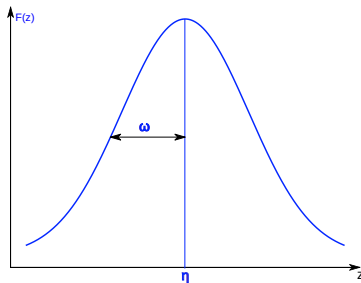
The individual-based model

Trait and fitness

Stabilizing selection: The **fitness** of an organism with trait value z is

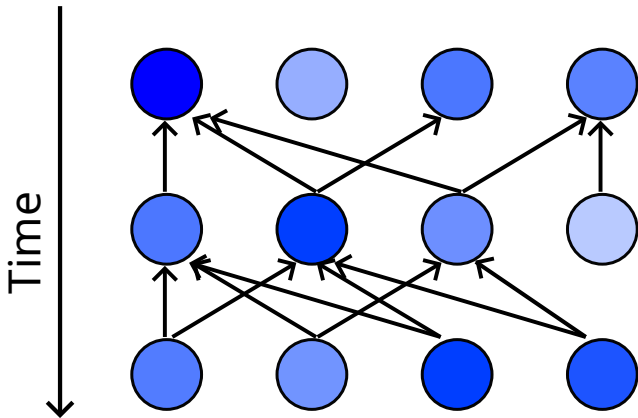
$$F(z) := \exp \left[-\frac{1}{2\omega^2} (z - \eta)^2 \right]$$

where ω^{-2} is the **strength of selection** and η is the **selection optimum**.



The individual-based model

Fitness and reproduction



Every generation, organisms pick two parents at random (\clubsuit) with probability proportional to fitness (\spadesuit). The population size is fixed at N .

Summary of the parameters

N	Population size
L	Number of loci
μ_ℓ^+, μ_ℓ^-	Mutation probabilities at locus ℓ
α_ℓ	Additive effect at locus ℓ
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This means $Z(G)$ is always in $[0, 2]$.

Three trait values

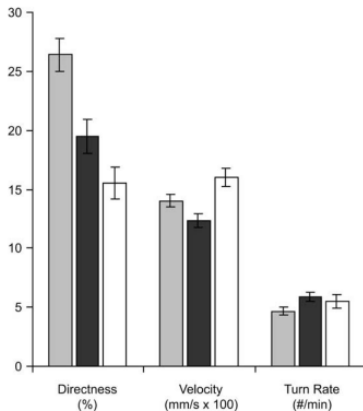
Three trait values of importance

The **selection optimum** is η .

The **heterozygote trait** is $\sum_{\ell \in [L]} \alpha_{\ell}$.

The **mutational optimum** is $\sum_{\ell \in [L]} 2\alpha_{\ell} \frac{\mu_{\ell}^{+}}{\mu_{\ell}^{+} + \mu_{\ell}^{-}}$.

On mutational bias



From Ajie et al 2005.

The roadmap

Individual-based
model
(L, N)



The polygenic limit

Features of the polygenic limit

There are three features which are expected to arise in the polygenic limit

- The trait is normally distributed (Quetelet 1835, Galton 1886, Fisher 1918), meaning if we know the population at present and sample a genome G uniformly at random, then $Z(G)$ is distributed as $\mathcal{N}(\bar{z}_t, \sigma^2)$.

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- The trait mean evolves as an Ornstein-Uhlenbeck process (Lande 1976)

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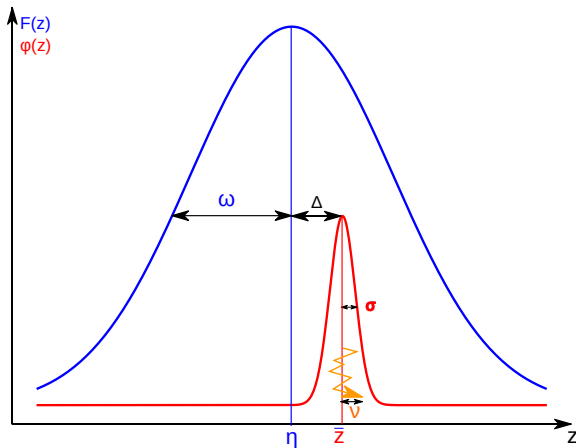
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- The dynamics at a locus can be described with an autonomous equation.

Macroscopic observables



The roadmap

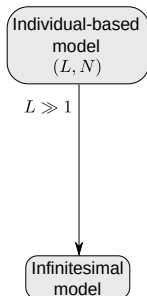
Individual-based
model
 (L, N)

Polygenic equation

Ornstein-Uhlenbeck
process

Trait normally
distributed

The roadmap



Polygenic equation

Ornstein-Uhlenbeck
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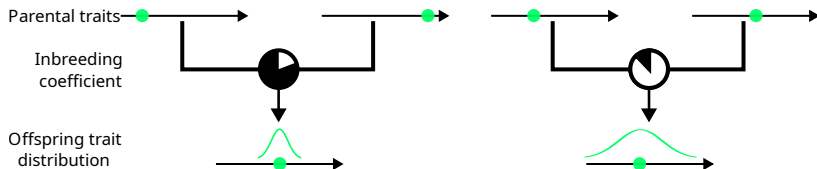
Trait normally
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The infinitesimal model

If we let $L \rightarrow +\infty$ in the previous model and scale the other parameters accordingly with N finite, we may obtain the infinitesimal model (Barton, Etheridge, Véber 2017).

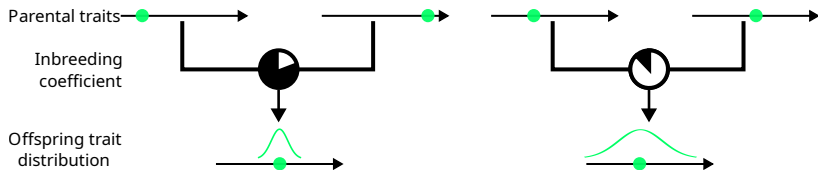
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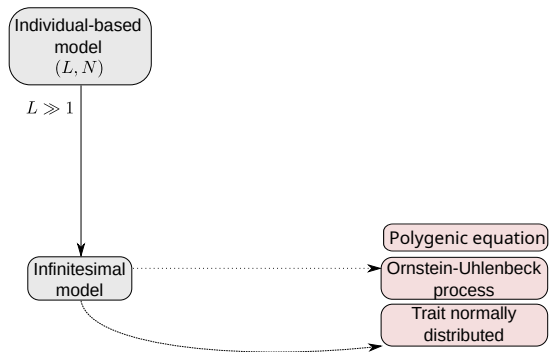
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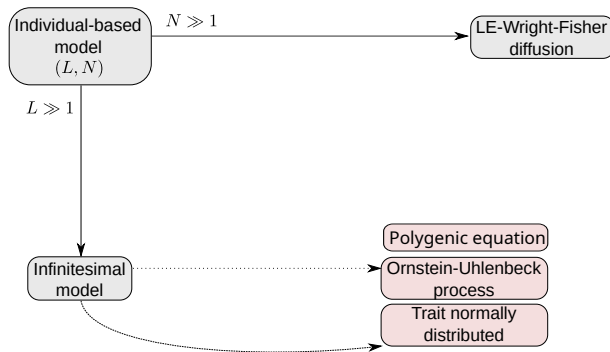


Key parameter: the segregation variance.

The roadmap

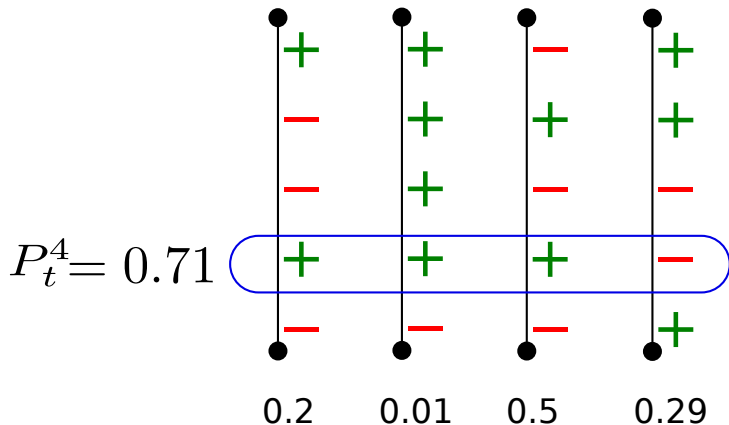


The roadmap



Taking the gene's eye-view

Let P_t^ℓ be the frequency of the $+$ allele at locus ℓ at generation $\lfloor 2Nt \rfloor$.



Taking the gene's eye-view

Step 0: rescaling

Let P_t^ℓ be the frequency of the $+$ allele at locus ℓ at generation $\lfloor 2Nt \rfloor$, with $\mathbf{P}_t = (P_t^\ell)_{\ell \in [L]}$.

We also define

$$\theta_\ell^\pm := 2N\mu_\ell^\pm$$

$$\omega_e^{-2} := 2N\omega^{-2}$$

Taking the gene's eye-view

Step 1: the diffusion approximation

First and second moment computations suggest the dynamics of $(P_t^\ell)_{t \geq 0}$ can be approximated with the SDE

$$dP_t^\ell = S^\ell(\mathbf{X}_t) \underset{\spadesuit}{P_t^\ell} (1 - P_t^\ell) dt + \left((1 - P_t^\ell) \underset{\diamond}{\theta_\ell^+} - P_t^\ell \underset{\clubsuit}{\theta_\ell^-} \right) dt + \sqrt{P_t^\ell (1 - P_t^\ell)} dB_t^\ell$$

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where we have (Robertson 1966)

$$S^\ell(\mathbf{X}_t) := 2N \frac{\mathbf{Cov}_{\mathbf{X}_t}[W(g), g_\ell]}{\mathbf{Var}_{\mathbf{X}_t}[g_\ell]}$$

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where $W(g) := \ln F(Z(g))$ is the **logfitness** and under \mathbf{X}_t , $g = (g_\ell)_{\ell \in [L]} \in \{0, 1, 2\}^L$ is the unphased genome of a randomly sampled organism at generation $\lfloor 2Nt \rfloor$ (this means $g_\ell = i$ if the organism has i copies of the $+$ allele at locus ℓ for $i \in [2]$).

Taking the gene's eye-view

Step 2: Hardy-Weinberg Linkage Equilibrium (HWLE)

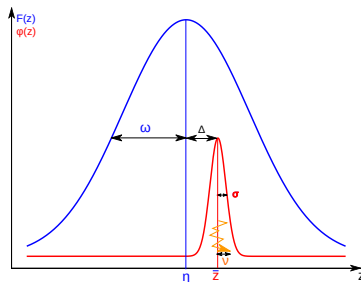
If recombination (♥) is strong then the population is close to HWLE, which means under \mathbf{X}_t , the $(g_\ell)_{\ell \in [L]}$ are independent $\text{Binomial}(2, P_t^\ell)$ variables.

Taking the gene's eye-view

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If recombination (♥) is strong then the population is close to HWLE, which means under \mathbf{X}_t , the $(g_\ell)_{\ell \in [L]}$ are independent $\text{Binomial}(2, P_t^\ell)$ variables. If we assume HWLE (Wright 1935) then we have

$$S^\ell(\mathbf{X}_t) = -\alpha_\ell \frac{\Delta_t}{\omega_e^2} + \frac{\alpha_\ell^2}{\omega_e^2} \left(P_t^\ell - \frac{1}{2} \right) =: \xi_{\alpha_\ell, \Delta_t}(P_t^\ell).$$



Taking the gene's eye-view

Step 3: LE-Wright-Fisher diffusion

We have

$$dP_t^\ell = \xi_{\Delta_t, \alpha_\ell}(\underset{\spadesuit}{P_t^\ell}) \underset{\diamond}{P_t^\ell} (1 - P_t^\ell) dt + ((1 - P_t^\ell) \theta^+ - \underset{\clubsuit}{P_t^\ell} \theta^-) dt + \sqrt{P_t^\ell (1 - P_t^\ell)} dB_t^\ell$$

with $\Delta_t = \sum_{\ell \in [L]} 2\alpha_\ell \underset{\spadesuit}{P_t^\ell} - \eta$.

Taking the gene's eye-view

Step 3: LE-Wright-Fisher diffusion

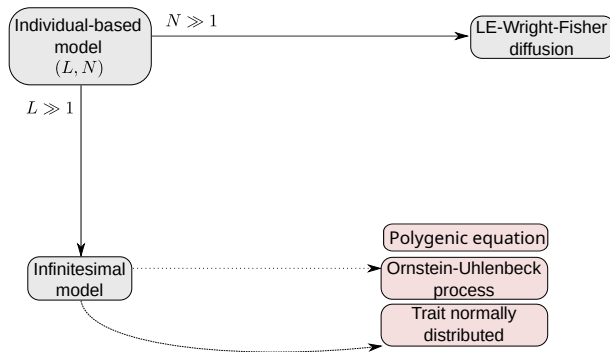
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with $\Delta_t = \sum_{\ell \in [L]} 2\alpha_\ell P_t^\ell - \eta$.

We call this the **LE-Wright-Fisher diffusion**.

The roadmap



Taking the gene's eye-view

Step 4: Propagation of chaos

Define the **typical locus** $(\vec{P}_t)_{t \geq 0} = (P_t, \alpha, \theta^\pm)_{t \geq 0}$ is defined as $(P_t^{\ell_U}, \alpha_{\ell_U}, \theta_{\ell_U}^\pm)$ where ℓ_U is picked uniformly at random on $[L]$.

Taking the gene's eye-view

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Taking the gene's eye-view

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We have

$$dP_t = \xi_{\Delta^*, \alpha}(P_t) P_t(1 - P_t) dt + ((1 - P_t)\theta^+ - P_t\theta^-) dt + \sqrt{P_t(1 - P_t)} dB_t \quad (1)$$

♠♦♣

If $L \gg 1$, we may replace Δ_t with

$$\Delta^* = \mathbb{E}[\Delta_t] = 2L\mathbb{E}[\alpha P_t] - \eta.$$

Taking the gene's eye-view

Step 5: fixed-point equation

If we know Δ^* , then P_t conditional on α, θ has stationary distribution $\Pi_{\Delta^*, \alpha, \theta}$ where for $\delta \in \mathbb{R}$

$$\Pi_{\delta, \alpha, \theta}(dp) \propto p^{2\theta^+ - 1} (1 - p)^{2\theta^- - 1} e^{\int_0^p \xi_{\delta, \alpha}(p') dp'} dp$$

Taking the gene's eye-view

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For $\delta \in \mathbb{R}$, define

$$I(\delta) := \mathbb{E} \left[\alpha \int p \Pi_{\delta, \alpha, \theta}(\mathrm{d}p) \right]$$

where the expectation is with respect to (α, θ) .

Taking the gene's eye-view

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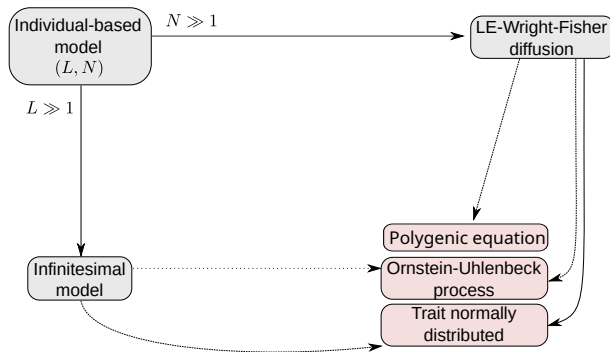
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Then

$$\Delta^* = 2LI(\Delta^*) - \eta.$$

(when $\Delta^* \ll 1$ and θ^+, θ^- are constant across loci and small, this equation was written and solved by Charlesworth 2013).

The roadmap



Orders of magnitude

L	Number of loci
$\theta_{\ell}^{+}, \theta_{\ell}^{-}$	Mutation rates at locus ℓ
α_{ℓ}	Additive effect at locus ℓ
ω_e^{-2}	Effective strength of selection
η	Selection optimum

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Define the mean mutation rate as $|\bar{\theta}| = \mathbb{E}[\theta^+ + \theta^-]$.

Assumptions

A 1: Uniform boundedness

There is a $C \asymp 1$ such that for any $\ell \in [L]$, $|\theta_\ell| \leq C|\bar{\theta}|$ and $\alpha_\ell \leq \frac{C}{L}$.

A 2: Mutations smaller than genetic drift $|\bar{\theta}| \lesssim 1$.

A 3: Mutational bias not too extreme

There is a constant $C \asymp 1$ such that for any ℓ , $\theta_\ell^-/C \leq \theta_\ell^+ \leq C\theta_\ell^-$.

A 4: Accessibility of the selection optimum

$\eta \in (0, 2)$ satisfies $\eta(2 - \eta) \asymp 1$.

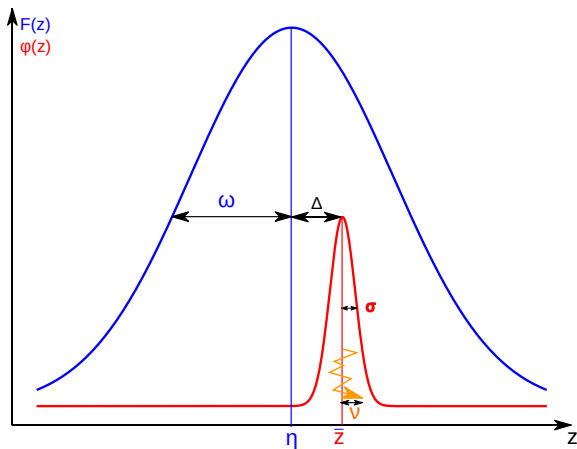
A 5: Weak/moderate/strong selection $L \lesssim \omega_e^{-2} \lesssim L^2$.

A 6: Distance between the selection and the mutation optimum

We have

$$|2L I(0) - \eta| \asymp 1.$$

Reminders



In the polygenic limit, under (A1-6), we have

$$|\Delta^*| \asymp L\omega_e^2$$

$$\sigma^2 \asymp \frac{|\bar{\theta}|}{L}$$

$$\nu^2 \asymp \omega_e^2$$

$$\rho \asymp \frac{|\bar{\theta}|}{L\omega_e^2}$$

The distance to the optimum is of the same order as its fluctuations

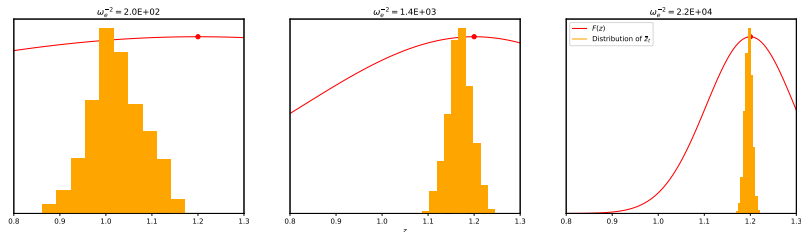


Figure: $N = 100, L = 100, \theta = (0.1, 0.2), \eta = 1.2$, and $(\alpha_\ell)_{\ell \in [L]}$ were sampled with distribution $\text{Exponential}(L)$

Recall the selection coefficient at the typical locus is

$$\xi_{\alpha, \Delta^*}(P_t) := -\alpha \frac{\Delta^*}{\omega_e^2} + \frac{\alpha^2}{\omega_e^2} \left(P_t - \frac{1}{2} \right)$$

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$$\xi_{\alpha, \Delta^*}(P_t) := L\alpha s^* + \frac{\alpha^2}{\omega_e^2} \left(P_t - \frac{1}{2} \right)$$

with

$$s^* := -\frac{\Delta_t}{L\omega_e^2} \asymp 1.$$

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with

$$s^* := -\frac{\Delta_t}{L\omega_e^2} \asymp 1.$$

When $L \ll \omega_e^{-2} \ll L^2$ (moderate selection), the value of s^* is independent of the strength of selection ω_e^{-2} .

Weak/moderate stabilizing selection has the same effect as directional selection

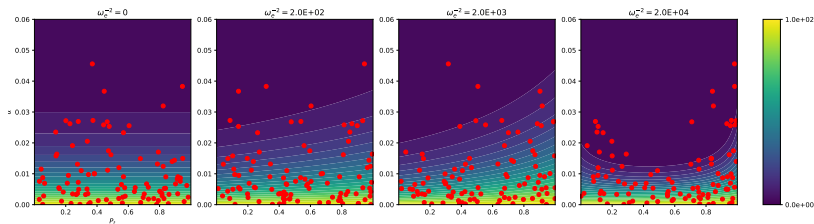


Figure: Parameters: $N = 500$, $L = 100$, $\theta = (0.5, 0.5)$, $\eta = 1.2$ and α_ℓ is distributed as $Exponential(L)$.

Macroscopic observables are well predicted by the fixed point equation

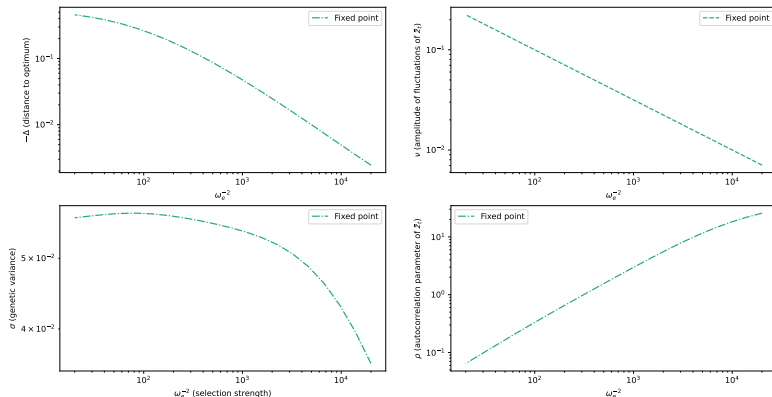


Figure: $L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 500N$ and α_ℓ is distributed as $\text{Exponential}(L)$.

Macroscopic observables are well predicted by the fixed point equation

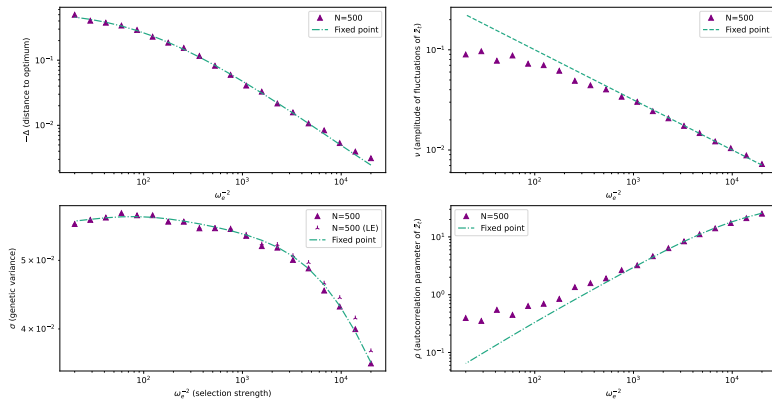


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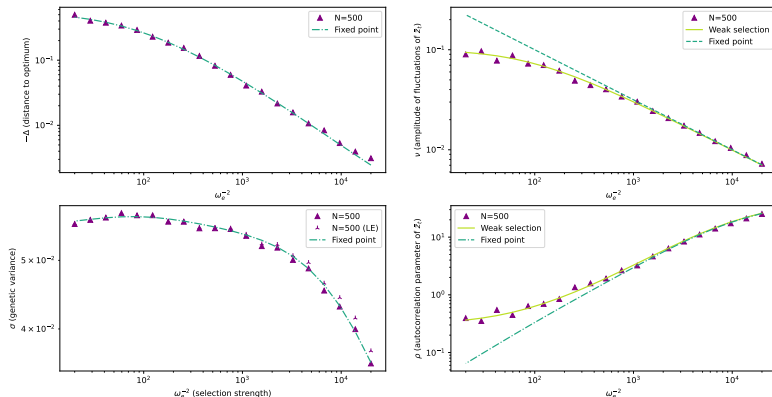


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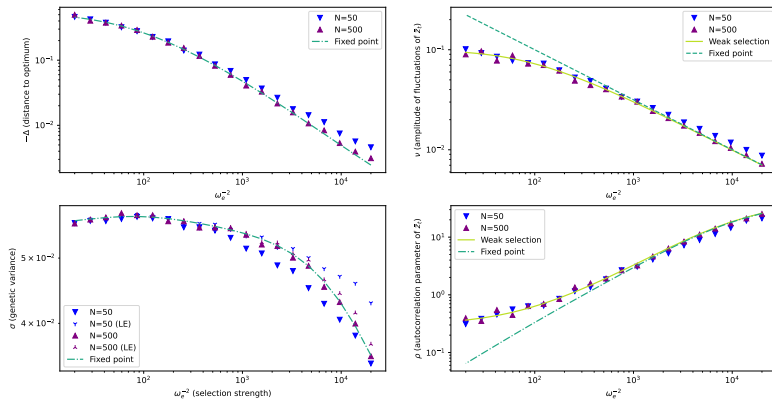
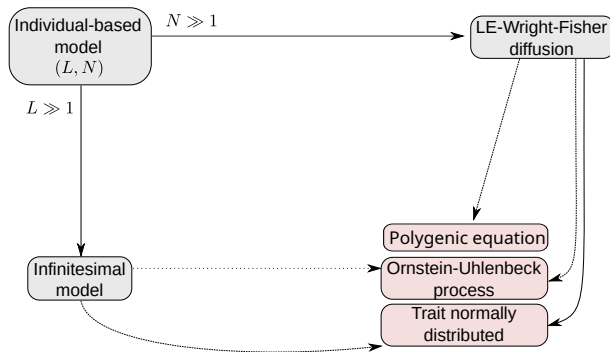
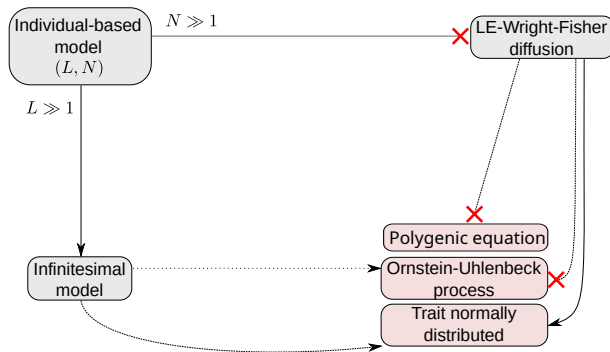


Figure: $L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 500N$ and α_ℓ is distributed as $\text{Exponential}(L)$.

The roadmap

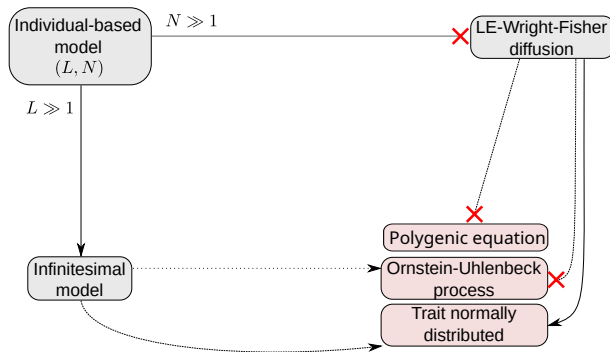


The roadmap



Breakdown of the polygenic limit

The roadmap



The Bulmer effect (1971)

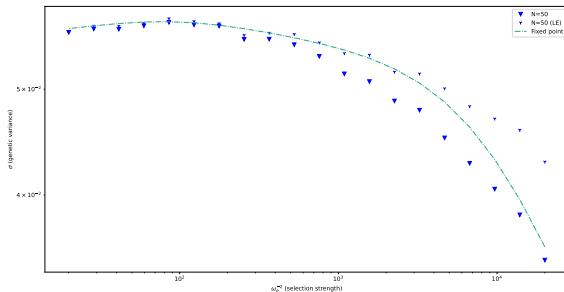


Figure: $L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 500N$ and α_ℓ is distributed as *Exponential*(L).

The Bulmer effect (1971)

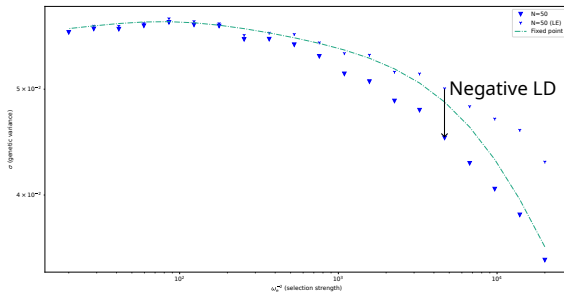


Figure: $L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 500N$ and α_{ℓ} is distributed as *Exponential*(L).

The Bulmer effect (1971)

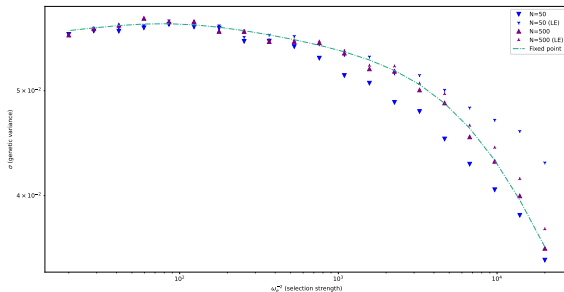


Figure: $L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 500N$ and α_ℓ is distributed as *Exponential*(L).

The Hill-Robertson effect (1966)

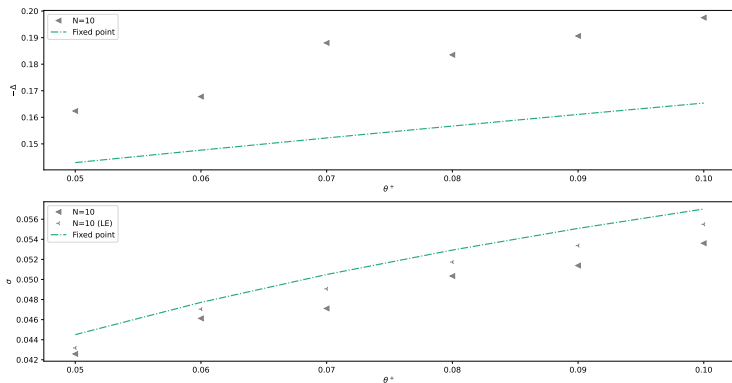


Figure: $L = 100$, $\theta^- = 2\theta^+$, $\eta = 1.2$, $\omega_e^{-2} = 2L$, $T = 500N$ and α_ℓ is distributed as $\text{Exponential}(L)$.

The Hill-Robertson effect (1966)

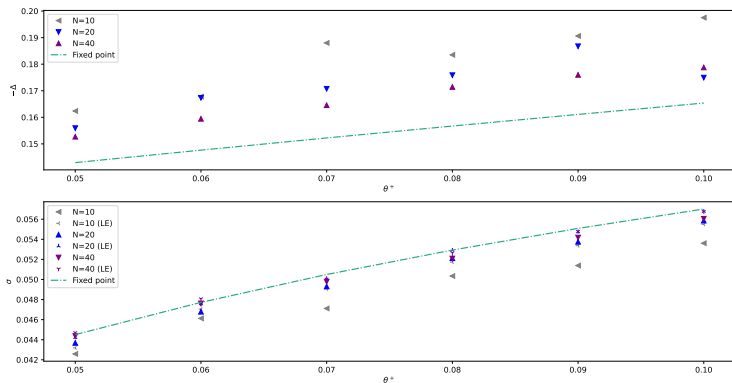


Figure: $L = 100, \theta^- = 2\theta^+, \eta = 1.2, \omega_e^{-2} = 2L, T = 500N$ and α_ℓ is distributed as $\text{Exponential}(L)$.

Assumptions

A 1: Uniform boundedness

There is a $C \asymp 1$ such that for any $\ell \in [L]$, $|\theta_\ell| \leq C|\bar{\theta}|$ and $\alpha_\ell \leq \frac{C}{L}$.

A 2: Mutations smaller than genetic drift $|\bar{\theta}| \lesssim 1$.

A 3: Mutational bias not too extreme

There is a constant $C \asymp 1$ such that for any ℓ , $\theta_\ell^-/C \leq \theta_\ell^+ \leq C\theta_\ell^-$.

A 4: Accessibility of the selection optimum

$\eta \in (0, 2)$ satisfies $\eta(2 - \eta) \asymp 1$.

A 5: Weak/moderate/strong selection $L \lesssim \omega_e^{-2} \lesssim L^2$.

A 6: Distance between the selection and the mutation optimum

We have

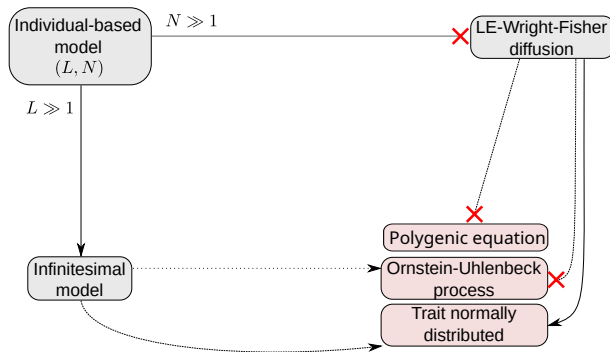
$$|2L I(0) - \eta| \asymp 1.$$

Under (A1-6), under the Quasi-Linkage Disequilibrium approach, the breakdown occurs when the following assumption fails

N1: sufficiently large population

$$2N \gg |\bar{\theta}| \frac{\ln(L)}{L\omega_e^2} + L \sqrt{|\bar{\theta}|}$$

The roadmap



Low mutations breakdown

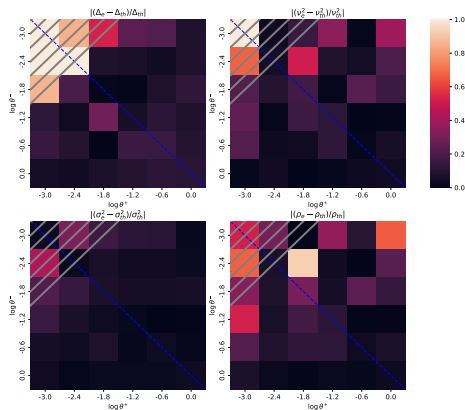


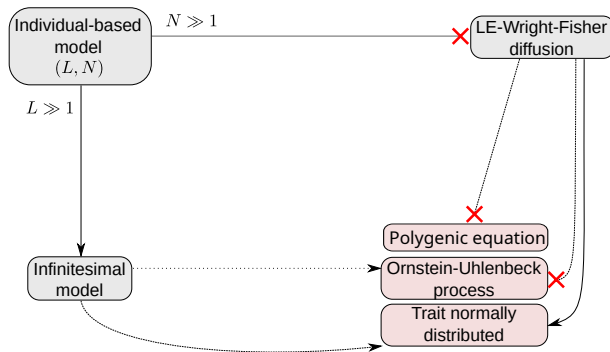
Figure: $N = 500, L = 100, \eta = 1.2, \omega_e^{-2} = 10^3, T = 500N$

Under (A1-6), the breakdown for Δ^* , σ^2 is expected when the following assumption fails

N2: Minimal mutational input every generation

$|\bar{\theta}|L \gg 1$.

The roadmap



The Ornstein-Uhlenbeck process is

$$d\Delta_t = \rho(\Delta^* - \Delta_t)dt + \nu \sqrt{2\rho} dB_t$$

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For such a process we expect

$$\ln(\mathbb{C}\mathbf{ov}[\Delta_t, \Delta_{t+u}]) = -\rho u + 2 \ln(\nu)$$

Ornstein-Uhlenbeck breakdown

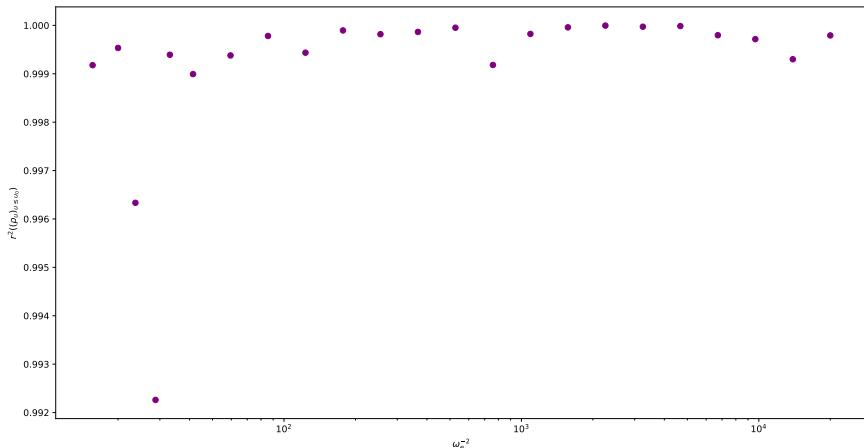


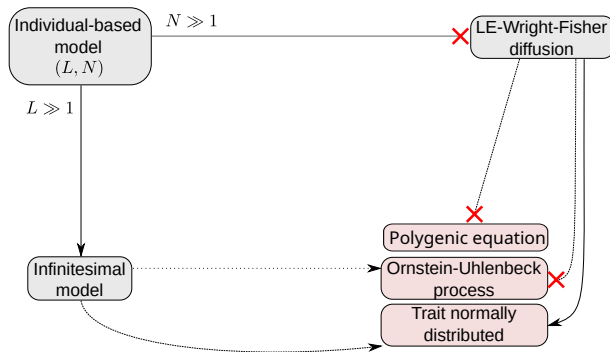
Figure: $N = 500, L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 1000N$

Under (A1-6), the breakdown for the description of $(\Delta_t)_{t \geq 0}$ as an Ornstein-Uhlenbeck process is expected when the following assumption fails

N3: Sufficient mutational input

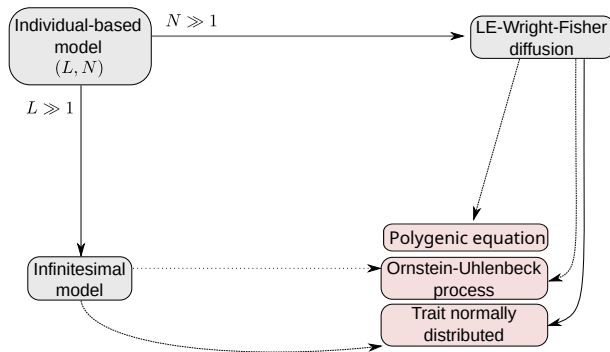
$$\frac{|\bar{\theta}|}{L\omega_e^2} \gg 1.$$

The roadmap

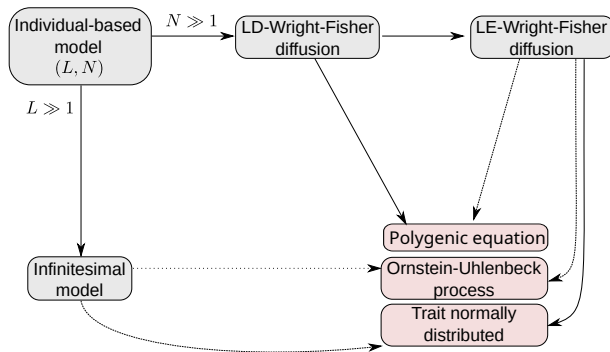


A rigorous control of LD

The roadmap



The roadmap



The LD-Wright-Fisher diffusion

We define $\mathbb{X}^{[L]}$ as the space of probability measures on $\{-1, +1\}^L$.

The LD-Wright-Fisher diffusion

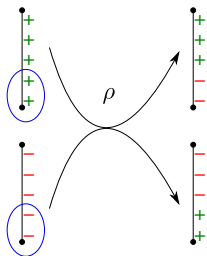
We define $\mathbb{X}^{[L]}$ as the space of probability measures on $\{-1, +1\}^L$.

We study a diffusion on $C([0, T], \mathbb{X}^{[L]})$ which includes all four forces of interest

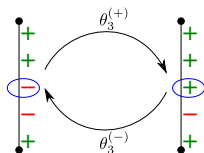
$$d\mathbf{X}_t = \underset{\heartsuit}{\rho R(\mathbf{X}_t)}dt + \underset{\spadesuit}{L S(\mathbf{X}_t)}dt + \underset{\diamondsuit}{\Theta(\mathbf{X}_t)}dt + \underset{\clubsuit}{\Sigma(\mathbf{X}_t)}d\mathbf{B}_t.$$

where $\rho > 0$ is the recombination rate and L represents the strength of selection.

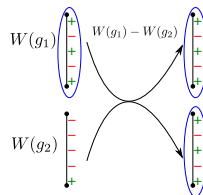
Operators



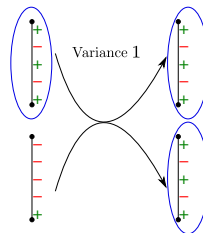
♥ Recombination



♦ Mutation



♠ Selection



♣ Genetic drift

Simplifications

- $\alpha_\ell = 1/L$

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Main theorem

Define the **local harmonic recombination rate** r_ℓ^* at locus ℓ with

$$\frac{1}{r_\ell^*} := \frac{1}{L-1} \sum_{\ell' \neq \ell} \frac{1}{r_{\{\ell, \ell'\}}}$$

where $r_{\{\ell, \ell'\}}$ is the probability of recombination between ℓ and ℓ' .

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Define the allelic law process as $\mu_{\mathbf{x}_t} := \frac{1}{L} \sum_{\ell \in [L]} \delta_{P_t^\ell}$.

Main theorem (C Lambert Schertzer 25+)

Theorem 1 (part 1)

Assume $\rho r^{**} \gg L^2 \ln(\rho)$. Then $(\mu_{\mathbf{x}_t})_{t \in [0, T]}$ converges to the law of the McKean-Vlasov process (1) in the Skorokhod topology.

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Theorem 1 (part 2)

Assume $\rho r^{**} \gg L^2 \ln(\rho)$. Consider n distinct loci $\ell_1 < \dots < \ell_n$ such that $(P_0^{\ell_i})_{i \in [n]}$ are independent and

$$\min_{i \in [n]} \rho r_{\ell_i}^* \gg L^2 \ln(\rho)$$

$$\min_{i, j \in [n]} \rho r_{\{\ell_i, \ell_j\}} \gg L$$

Then $(P_t^{\ell_i})_{i \in [n], t \in [0, T]}$ converge to n independent McKean-Vlasov processes (1).

Theorem 2

Assume $\rho r^{**} \gg L^2 \ln(\rho)$. Set

$$\varepsilon_L := \frac{1}{\sqrt{\rho r^{**}}}$$

Define the genetic variance $\sigma_t^2 := 4\mathbb{E}[f_t(1 - f_t)]$. Then

$$\mathbb{E} \left[\sup_{t \in [\varepsilon_L, T]} |L \mathbf{Var}_{\mathbf{x}_t}[Z(g)] - \sigma_t^2| \right] \rightarrow 0$$

Structure of the proof

For $\mathbf{x} \in \mathbb{X}^{[L]}$, let $\pi(\mathbf{x})$ be the **LE projection** of \mathbf{x}

$$\pi(\mathbf{x}) = \bigotimes_{\ell \in [L]} \mathbf{x}^{\{\ell\}}.$$

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- We can decompose the marginal of the selector on locus ℓ , $S^\ell(\mathbf{X}_t)$, into the LE term $S^\ell(\pi(\mathbf{X}_t))$ and a sum over subsets $A \subseteq [L]$, $\#A \leq 3$ of a term dominated by $\mathbf{X}_t^A - \pi(\mathbf{X}_t^A)$.

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Structure of the proof

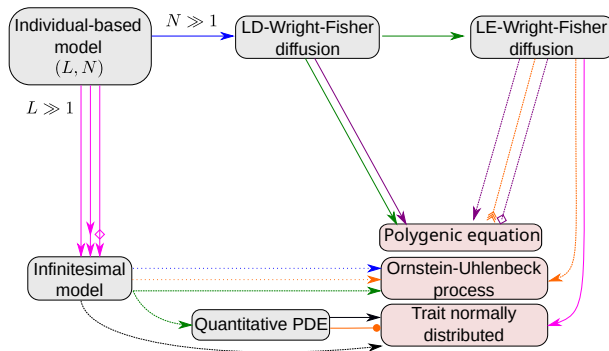
For $\mathbf{x} \in \mathbb{X}^{[L]}$, let $\pi(\mathbf{x})$ be the **LE projection** of \mathbf{x}

$$\pi(\mathbf{x}) = \bigotimes_{\ell \in [L]} \mathbf{x}^{\{\ell\}}.$$

- We can decompose the marginal of the selector on locus ℓ , $S^\ell(\mathbf{X}_t)$, into the LE term $S^\ell(\pi(\mathbf{X}_t))$ and a sum over subsets $A \subseteq [L]$, $\#A \leq 3$ of a term dominated by $\mathbf{X}_t^A - \pi(\mathbf{X}_t^A)$.
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- We can then show the convergence of the martingale problem associated with P_t^ℓ to the limit martingale problem.
- The well-posedness of the limit martingale problem is obtained from a Girsanov transform.

Conclusion

Conclusion for mathematicians



Degree of proof

- Proven
- Discussed
-→ Not discussed

Type of proof

- Diffusion approximation
- Mixing
- Mean-field approximation
- Separation of timescales
- Central limit theorem

Complications

- Diploidy/Dominance
- Strong selection
- Epistasis
- Population structure/Demography

Conclusion for mathematicians

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Conclusion for biologists

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- When genetic diversity is maintained by mutations, the demographic noise in trait mean should be of the same order as the distance of the trait mean to the selection optimum: $|\Delta_t| \gtrsim v$.

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Conclusion for biologists

- When genetic diversity is maintained by mutations, the demographic noise in trait mean should be of the same order as the distance of the trait mean to the selection optimum: $|\Delta_t| \gtrsim \nu$.
- We discuss how the gene's eye-view can be extended to account for polyploidy, pleiotropy, dominance, and some forms of epistasis. We can also hope to extend it to spatial structure, demographic variability and more.

Colleague acknowledgements

Supervisors: Amaury Lambert, Emmanuel Schertzer

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And also: Melda Akyazi, Nick Barton, Nathanaël Berestycki, Reinhard Bürger, Ariane Carrance, Archana Devi, Laura Hayward, Jitka Polechova, Himani Sachdeva, William da Silva, Lucas Teyssier, Paul Thévenin.

Appendix

First-order Bulmer correction

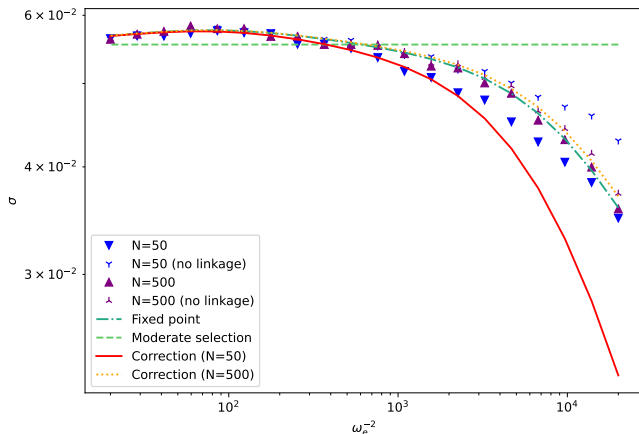
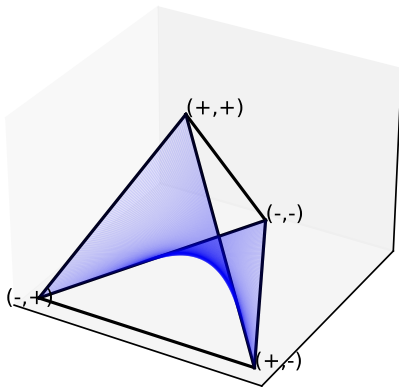


Figure: $L = 100, \theta = (0.1, 0.2), \eta = 1.2, T = 500N$ and α_ℓ is distributed as *Exponential*(L).



Slow/fast principle for strong selection

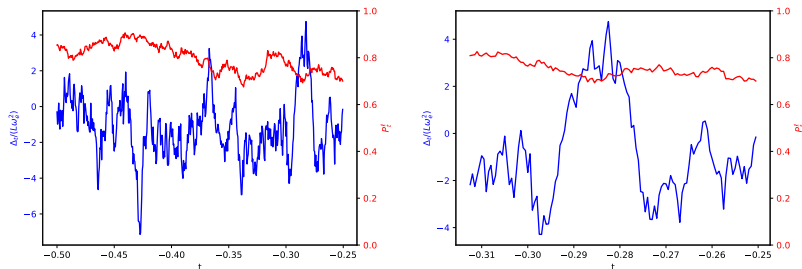


Figure: $N = 1000, L = 1000, \theta = (0.1, 0.2), \eta = 1.2, \omega_e^{-2} = L^2$ and α_ℓ has law $\text{Exponential}(L)$

Breakdown for small mutations.

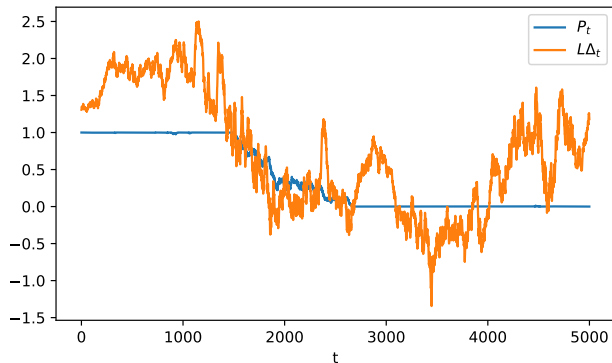
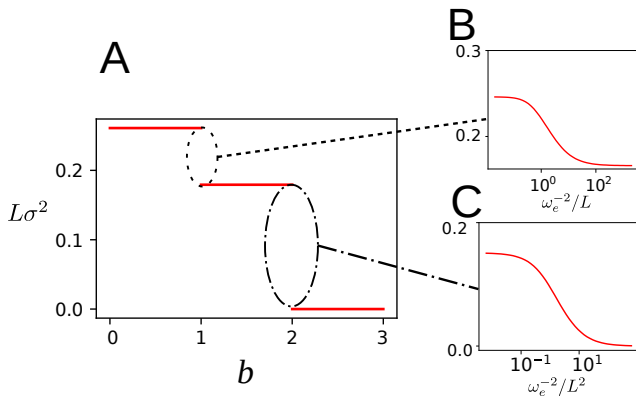


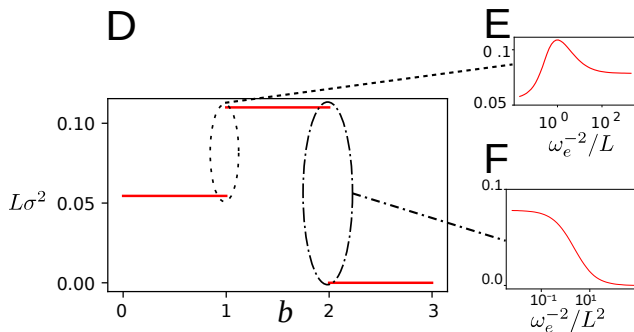
Figure: $N = 500, L = 100, \theta = (1/(2L), 1/(2L)), \omega_e^{-2} = 2000, \eta = 1.2$, and α_ℓ has law *Exponential*(L).

Scaling limit



with b such that $\omega_e^{-2} = L^b$, $L \rightarrow +\infty$, $N \rightarrow +\infty$, $\eta = 1.5$, and $(\alpha_\ell)_{\ell \in [L]}$ be exponentially distributed with parameter L .

No mutational bias ($\theta = (0.1, 0.1)$)



with b such that $\omega_e^{-2} = L^b$, $L \rightarrow +\infty$, $N \rightarrow +\infty$, $\eta = 1.5$, and $(\alpha_\ell)_{\ell \in [L]}$ be exponentially distributed with parameter L .

Strong mutational bias ($\theta = (0.01, 0.1)$)

Trajectory view of the Hill-Robertson effect

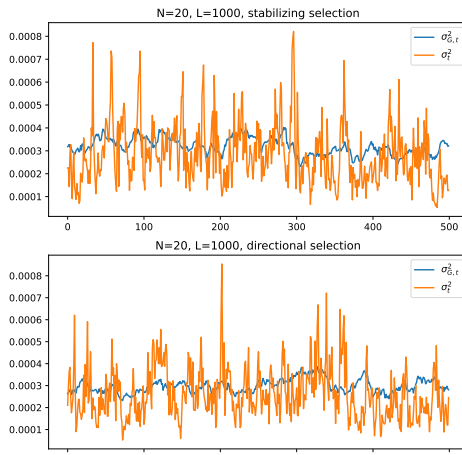


Figure: $\theta = (0.1, 0.2), \eta = 1.2, \omega_e^{-2} = 2L$.

Heavy tails of α

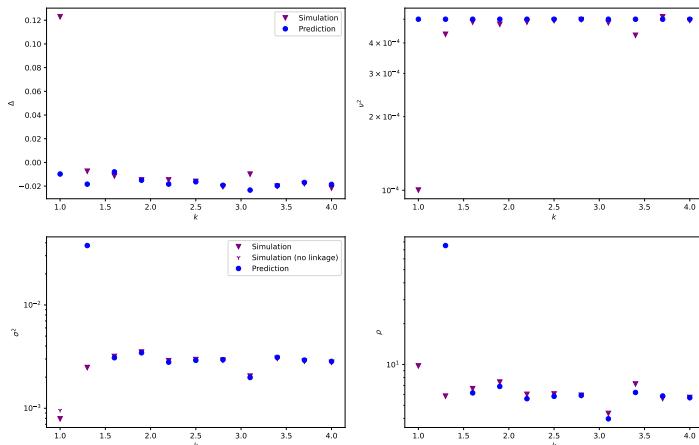


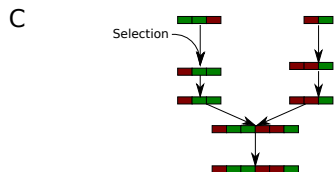
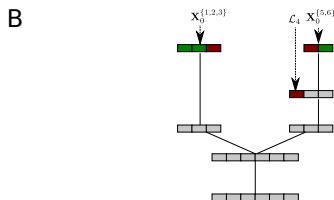
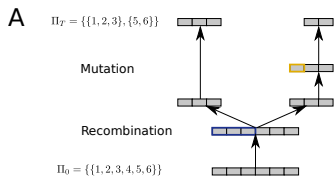
Figure: $N = 500, L = 100, \theta = (0.1, 0.2), \eta = 1.2, \omega_e^{-2} = 10^3$ and α has law $Pareto(k)$.

An ancestral process in an infinite population

$$\frac{d}{dt}\mathbf{X}_t = R(\mathbf{X}_t) + S(\mathbf{X}_t) + \Theta(\mathbf{X}_t)$$

♥ ♠ ♦

An ancestral process in an infinite population



Duality formula

For $\mathbf{x} \in \mathbb{X}^{[L]}$, $\pi \in \mathcal{S}$, $\gamma \in \square_{[L]}$ define

$$f(\mathbf{x}, \pi, \gamma) := \prod_{A \in \pi} x^A(\gamma^{|A|}) \times \prod_{\ell \notin \pi} \mathcal{L}_\ell(\gamma^{\{|\ell|\}}).$$

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We have

$$\mathcal{A}^{\mathbf{x}_t} f(\cdot, \pi, \gamma)(\mathbf{x}) = \mathcal{A}^{\Pi_t} f(\mathbf{x}, \cdot, \gamma)(\pi) + \mathcal{A}_{\mathbf{x}}^{\mathcal{G}_t} f(\mathbf{x}, \pi, \cdot)(\gamma).$$