

# Convergence of bivariate Markov chains to multi-type self-similar processes, and applications to scaling limits of some random trees

Robin Stephenson  
Universität Zürich

Joint work with Bénédicte Haas

# Introduction

# The model

We consider a discrete-time Markov chain  $(X, J)$ , taking values in  $\mathbb{Z}_+ \times \{1, \dots, \kappa\}$ , such that the  $X$  component is **nonincreasing**.

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Notation :  $\left( (X_n^{(i)}(k), J_n^{(i)}(k)), k \geq 0 \right)$  is a version of the process starting at  $(n, i)$ .

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Basic principle : assume that the macroscopic jumps are rare, *i.e.* that there exists  $\gamma > 0$  such that

$$\mathbb{P}[X_n^{(i)}(1) \leq (1 - \varepsilon)n] \underset{n \rightarrow \infty}{\sim} c_\varepsilon^{(i)} n^{-\gamma}, \quad \forall \varepsilon > 0, \forall i \in \{1, \dots, \kappa\}$$

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Then

$$\left( \frac{X_n^{(i)}(\lfloor n^\gamma t \rfloor)}{n}, t \geq 0 \right) \xrightarrow{(d)} (X_\infty^{(i)}(t), t \geq 0)$$

where  $X_\infty^{(i)}$  is some kind of self-similar Markov process.

## Previous work

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- Haas & Miermont (2011) : the monotype case ( $K = 1$ ).
- Bertoin & Kortchemski (2015) : still monotype case, but  $X$  is not assumed to be nonincreasing.

## The monotype case

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- $\alpha$  is a real parameter called the *index of self-similarity*.

# The result

## Theorem (Haas-Miermont)

Assume that, for any continuous function  $f$  on  $[0, 1]$ ,

$$n^\gamma \mathbb{E} \left[ \left( 1 - \frac{X_n(1)}{n} \right) f \left( \frac{X_n(1)}{n} \right) \right] \xrightarrow[n \rightarrow \infty]{(d)} \int_{[0,1]} f(x) d\mu(x),$$

where  $\gamma > 0$  and  $\mu$  is a finite and nontrivial measure on  $[0, 1]$ .

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Then

$$\left( \frac{X_n(\lfloor n^\gamma t \rfloor)}{n}, t \geq 0 \right) \xrightarrow{(d)} (X_\infty(t), t \geq 0)$$

where  $X_\infty$  is a pssMp with index of self-similarity  $\gamma$ , and the underlying Lévy process  $\xi$  is a subordinator with Laplace exponent  $\psi$  such that

$$\psi(\lambda) = \mu(\{0\}) + \mu(\{1\})\lambda + \int_{(0,1)} (1 - x^\lambda) \frac{d\mu(x)}{1 - x}.$$

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The underlying topology is the Skorokhod topology on the space of càdlàg functions on  $[0, \infty)$ .

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$$A_n = \inf\{k \in \mathbb{N} : \forall l, X_n(k+l) = X_n(k)\}$$

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then

$$\left(\frac{X_n(\lfloor n^\gamma \cdot \rfloor)}{n}, \frac{A_n}{n}\right) \xrightarrow{(d)} (X_\infty(\cdot), \sigma).$$

## The limiting processes : Markov Additive Processes and their Lamperti transforms

# Markov Additive Processes (MAP)

Let  $((\xi_t, K_t), t \geq 0)$  be a Markov process on  $\mathbb{R} \times \{1, \dots, \kappa\}$ , such that the  $\xi$  component is nondecreasing. We write  $\mathbb{P}_{(x,i)}$  for its distribution when starting at a point  $(x, i)$ ,

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$$\left( (\xi_{t+s} - \xi_t, K_{t+s}), s \geq 0 \right) \text{ has distribution } \mathbb{P}_{0, K_t}.$$

## In practice - parametrisation

We think of  $(\xi, K)$  as a "typed subordinator" :

- $K$  is a continuous-time Markov chain on  $\{1, \dots, \kappa\}$ , with transition rates  $(\lambda_{i,j})_{(i,j) \in \{1, \dots, \kappa\}^2, i \neq j}$ .

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- On the intervals of constancy of  $K$ ,  $\xi$  is a subordinator with Laplace exponent  $\psi_i$ .
- Each jump of  $K$  induces a jump of  $\xi$ , and we call  $B_{i,j}$  its distribution if we jump from  $i$  to  $j \neq i$ .

# Lamperti transforms

Let  $\alpha \in \mathbb{R}$ , we call the Lamperti transform of  $(\xi, K)$  the process  $(X, J)$  defined by

$$X_t = e^{-\xi_{\tau(t)}}, \quad J_t = K_{\tau(t)},$$

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Note that, if  $\alpha > 0$ , the death time  $\sigma = \inf\{t \geq 0 : X(t) = 0\} = \int_0^\infty e^{-\alpha \xi_s} ds$  is finite.  $(X, J)$  is càdlàg on  $[0, \sigma)$ , but  $J$  does not have a limit at  $\sigma$ .

## Main results

## Three regimes

The nature of the limiting process depends on the rate of change of the type  $J$  : assume that there is some  $\beta \geq 0$  such that

$$\forall j \neq i, \mathbb{P}[J_n^{(i)}(1) = j] \sim p_{i,j} n^{-\beta}.$$

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- Mixing regime : if  $0 \leq \beta < \gamma$  then the types "mix" and disappear in the limit,  $X_\infty^{(i)}$  is then a pssMp which doesn't depend on  $i$ .

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- Mixing regime : if  $0 \leq \beta < \gamma$  then the types "mix" and disappear in the limit,  $X_\infty^{(i)}$  is then a pssMp which doesn't depend on  $i$ .
- Solo regime : if  $\beta > \gamma$ , the type does not change in the limit, and  $X_\infty^{(i)}$  is a pssMp which depends on  $i$ .

## Critical regime : assumption ( $H_{cr}$ )

Assume that, for all  $i, j \in \{1, \dots, \kappa\}$ , there exists finite measures  $\mu^{(i,j)}$  on  $(0, 1]$ , such that for all continuous functions  $f : [0, 1] \rightarrow \mathbb{R}$ ,

$$n^\gamma \mathbb{E} \left[ \left( 1 - \frac{X_n^{(i)}(1)}{n} \mathbf{1}_{\{j=i\}} \right) f \left( \frac{X_n^{(i)}(1)}{n} \right) \right] \xrightarrow[n \rightarrow \infty]{(d)} \int_{[0,1]} f(x) d\mu^{(i,j)}(x),$$

Moreover, for all  $i \in \{1, \dots, \kappa\}$ , at least one of the measure  $\mu^{(i,j)}$ ,  $j \in \{1, \dots, \kappa\}$  is not trivial.

## Critical regime : main convergence

### Theorem (Haas-S.)

Assume  $(H_{\text{cr}})$ . Then, for all  $i \in \{1, \dots, \kappa\}$ ,

$$\left( \frac{X_n^{(i)}(\lfloor n^\gamma \cdot \rfloor)}{n} \right) \xrightarrow{(d)} X_\infty^{(i)}(\cdot),$$

where  $X_\infty^{(i)}$  is the first component of a Lamperti MAP with the following parameters :

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- The self-similarity index is  $\gamma$ .
- $\psi_i(\lambda) = \mu^{(i,i)}(\{0\}) + \mu^{(i,i)}(\{1\})\lambda + \int_{(0,1)} (1 - x^\lambda) \frac{d\mu^{(i,i)}(x)}{1-x}$ .
- $\lambda_{i,j} B_{i,j} = \mu^{(i,j)} \circ (-\log)^{-1}$ .
- The initial type is  $i$ .

## Critical regime : death times

### Theorem (Haas-S.)

*Assume moreover that, for all  $i \in \{1, \dots, \kappa\}$ , there exists  $j$  such that  $\mu^{(i,j)}([0, 1)) > 0$ .*

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Then, calling  $A_n^{(i)}$  the death time of  $X_n^{(i)}$  and  $\sigma^{(i)}$  that of  $X_\infty^{(i)}$ , we have, jointly with the previous convergence,

$$\frac{A_n^{(i)}}{n} \xrightarrow{(d)} \sigma^{(i)}.$$

## Mixing regime : assumption ( $H_{\text{mix}}$ )

Assume that there exists  $\beta \geq 0$  with  $\beta < \gamma$  for which :

- (i) There exist finite measures  $(\mu^{(i)}, i \in \{1, \dots, \kappa\})$  on  $[0, 1]$ , such that, for all continuous functions  $f : [0, 1] \rightarrow \mathbb{R}$ ,

$$n^\gamma \mathbb{E} \left[ f \left( \frac{X_n^{(i)}(1)}{n} \right) \left( 1 - \frac{X_n^{(i)}(1)}{n} \right) \right] \xrightarrow[n \rightarrow \infty]{} \int_{[0,1]} f(x) d\mu^{(i)}(x).$$

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- (ii) Moreover, there exists a  $Q$ -matrix  $Q = (q_{i,j})_{i,j \in \{1, \dots, \kappa\}}$  having a unique irreducible component, such that, for all types  $i \neq j$

$$n^\beta \mathbb{P}[J_n^{(i)}(1) = j] \underset{n \rightarrow \infty}{\sim} n^{-\beta} q_{i,j}$$

and

$$n^\beta (\mathbb{P}[J_n^{(i)}(1) = i] - 1) \underset{n \rightarrow \infty}{\sim} n^{-\beta} q_{i,i}.$$

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We call  $\pi = (\pi_i)_{i \in \{1, \dots, \kappa\}}$  the irreducible distribution on  $\{1, \dots, \kappa\}$  associated to  $Q$ .

## Mixing regime : main convergence

### Theorem (Haas-S.)

Assume  $(H_{\text{mix}})$ . Then, for all  $i \in \{1, \dots, \kappa\}$ ,

$$\left( \frac{X_n^{(i)}(\lfloor n^\gamma \cdot \rfloor)}{n} \right) \xrightarrow{(d)} (X_\infty(\cdot)),$$

where  $X_\infty$  is a pssMp with the following parameters :

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- The self-similarity index is  $\gamma$ .

- $\psi(\lambda) =$

$$\sum_{i=1}^{\kappa} \pi_i \left( \mu^{(i)}(\{0\}) + \mu^{(i)}(\{1\})\lambda + \int_{(0,1)} (1-x^\lambda) \frac{d\mu^{(i)}(x)}{1-x} \right).$$

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*Then, with similar notation to earlier, jointly with the previous convergence,*

$$\frac{A_n^{(i)}}{n} \xrightarrow{(d)} \sigma.$$

## Solo regime : assumption ( $H_{\text{sol}}$ )

We fix a type  $i$  and assume the following :

- (i) There exists a nontrivial finite measure  $\mu^{(i)}$  on  $[0, 1]$ , such that, for all continuous functions  $f : [0, 1] \rightarrow \mathbb{R}$ ,

$$n^\gamma \mathbb{E} \left[ f \left( \frac{X_n^{(i)}(1)}{n} \right) \left( 1 - \frac{X_n^{(i)}(1)}{n} \right) \right] \xrightarrow{n \rightarrow \infty} \int_{[0,1]} f(x) d\mu^{(i)}(x).$$

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- (ii) Moreover,

$$n^\gamma \mathbb{P}(J_n^{(i)}(1) \neq i) \xrightarrow{n \rightarrow \infty} 0.$$

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Assume moreover to  $(H_{\text{sol}})$  that, for some  $a < 1$  and for all types  $j$ ,

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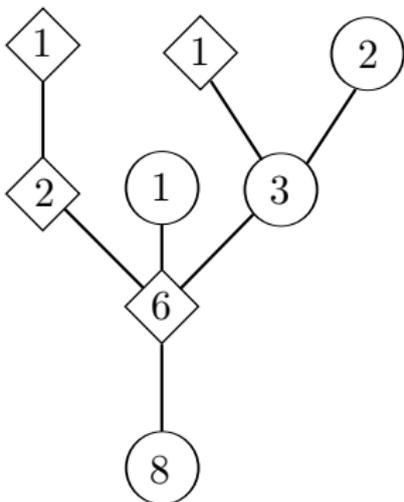
## Multi-type Markov branching trees

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# Definition

- Offspring distribution : for all  $n$  and  $i$ ,  $q_n^{(i)}$  is the distribution of the offspring of an individual of size  $n$  and type  $i$ . It is a probability measure on

$$\bar{\mathcal{P}}_n = \left\{ ((n_1, t_1), \dots, (n_k, t_k)) : n_1 \geq n_2 \geq n_k, \sum n_i \leq n, t_i \in \{1, \dots, \kappa\} \right\}$$

- We call  $T_n^{(i)}$  a version of the tree with an ancestor of size  $n$  and type  $i$
- Sometimes it is convenient for the tree to be *planted* : an extra untyped vertex is added below the root
- The monotype trees were studied by Haas and Miermont (2012)

## Example 1 : Contioned multi-type Galton-Watson trees

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- the number of vertices of fixed type  $i$ . The size of a vertex is then the number of vertices of type  $i$  in the descending subtree.
- more generally, we can condition on  $\sum_{i=1}^{\kappa} \alpha_i |T|_i$ , where  $|T|_i$  is the number of vertices of type  $i$ ...

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The case where  $S$  is a star is the subject of previous work (2015).

## Example 2 : Recursively growing trees

This yields a (planted) Markov branching tree if :

- Each vertex of  $S$  has a type, with 1 for its root.
- The size of a vertex is the number of vertices of type 1 in its descending subtree.

(with eventual "superfluous" types which we can identify if we want to.)

The scaling limits of multi-type Markov  
branching trees : multi-type self-similar  
fragmentation trees

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Call  $X_n$  the biggest element of the first generation of a  $T_n$ . Assume that

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$$\frac{1}{n^\gamma} T_n \xrightarrow{(d)} \mathcal{T}$$

where  $\mathcal{T}$  is a *self-similar fragmentation tree* with explicit distribution.

# Multi-type self-similar fragmentation processes and trees

Let, for  $i \in \{1, \dots, \kappa\}$ ,  $\nu_i$  be a  $\sigma$ -finite measure on

$$\mathcal{S}^\downarrow = \left\{ \mathbf{s} = (s_i, t_i)_{i \in \mathbb{N}} : s_1 \geq s_2 \geq \dots \geq 0, \sum s_i \leq 1, t_i \in \{1, \dots, \kappa\} \right\}$$

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A multi-type self-similar fragmentation process with self-similarity index  $\alpha \in \mathbb{R}$  and *dislocation measures*  $(\nu_i)$  is a  $\mathcal{S}^\downarrow$ -valued process such that a particle  $(x, i)$  transforms into a set of particles with masses and types  $x\mathbf{s} = (xs_i, t_i)_{i \in \mathbb{N}}$  at rate  $x^\alpha d\nu_i(\mathbf{s})$ .

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When  $\alpha < 0$ , a multi-type self-similar fragmentation *tree* with the same parameters is, informally, the family tree of the above process.

## Upcoming theorem

Let  $X_n^{(i)}$  be the largest element of the first generation of  $T_n^{(i)}$ , and  $J_n(i)$  its type. Assume that, for some  $\gamma > 0$  and  $\beta \geq 0$ ,

$$\mathbb{P}[X_n^{(i)} \leq (1 - \varepsilon)n] \underset{n \rightarrow \infty}{\sim} c_\varepsilon^{(i)} n^{-\gamma}, \quad \forall \varepsilon > 0, i \in \{1, \dots, \kappa\}$$

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for the Gromov-Hausdorff-Prokhorov topology, where  $\mathcal{T}^{(i)}$  is a fragmentation tree with self-similarity index  $-\gamma$  and explicit dislocation measures, and which is :

- multi-type if  $\beta = \gamma$ .
- monotype, not depending on  $i$  if  $0 \leq \beta < \gamma$ .
- monotype, depending on  $i$  if  $\beta > \gamma$ .

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- Our method will let us obtain this convergence for more general conditionings if the offspring distributions have exponential moments
- We will also obtain convergence to the stable trees, at least when conditioning on the number of vertices of one type.
- For the recursively growing trees,  $n^{-1/|S|}T_n$  converges in distribution, and maybe in probability, to a multi-type fragmentation tree.

Thank you for your attention